

Workshop on Dichotomous Rasch Modeling with Ministep/Winsteps

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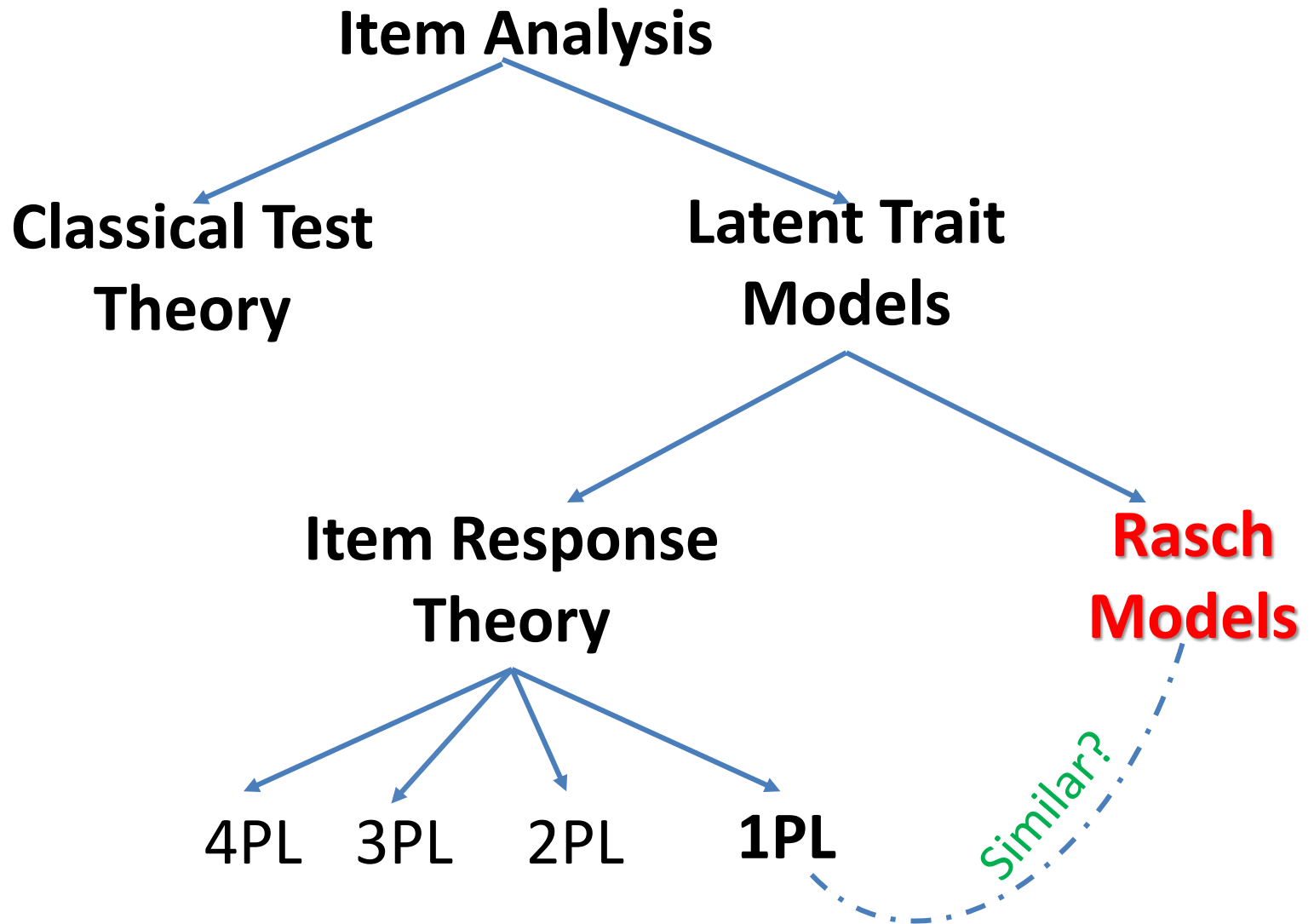
PEMEA – Continuing Professional Development Program

DLS-CSB Auditorium

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Shortcomings of CTT

1. Examinee's ability is **exam-dependent**

For a fixed length test, examinee's ability is high if the test is easy; and examinee's ability is low if the test is difficult. Therefore, examinee's ability is exam-dependent.

2. Item/Test Difficulty is **group-dependent**

Item is easy if higher abilities take the test; and item is difficult if low abilities take the test. Therefore, item/test difficulty is group-dependent.

3. CTT is test-oriented.

Score is given at the test level, but there is no basis in determining how well an examinee perform a particular item.

Limitations of CTT

1. Cannot predict (probabilistically) an examinee's response to an item.
2. Cannot predict individuals performance on certain items unless items have been administered to similar (comparable) individuals.
3. In adaptive testing, no mechanism exists in determining which item (*from an item pool*) is most appropriate to administer next.
4. Cannot determine how effective an item is at each level of ability.
5. Cannot estimate an examinee's ability from any given set of items

CTT Approach for Item Analysis using CITAS


CITAS is FREE!

CITAS - Free Version - Excel (Product Activation Failed)

FILE HOME INSERT PAGE LAYOUT FORMULAS DATA REVIEW VIEW

Clipboard Font Alignment Number Styles Cells

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| | A | B | C | D | E | F | G | H | I | J | K | L | M |
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| 1 |  | | Assessment Systems Corporation | | | | | | | | | | |
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www.assess.com

Instructions for the Classical Item and Test Analysis Spreadsheet (CITAS)

1. Paste your data (ABCDE) in the blue data range on the "Input" tab
2. Paste your keys (ABCDE) in the blue keys range in the top row of the "Input" tab
3. Output will automatically populate on the "Output" and "Response counts" tabs

Note: The "Scored 0-1" tab is calculations only, and is locked to preserve accuracy

CITAS is limited to 50 items and 50 people. For larger data sets and more functionality, please visit the webpages for Lertap and Iteman.

Instructions | Input | Scored 0-1 | Output | Response counts

Rasch Model

$$P(X_{vi} = 1 | \beta_v, D_i) = P_{vi} = \frac{e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}}$$

$$P(X_{vi} = 0 | \beta_v, D_i) = 1 - P_{vi}$$

P_{vi} = probability that person v gets a correct answer on item i , given his/her ability β and item difficulty D .

β_v = ability of examinee

D_i = difficulty of item i

$e = 2.718$ (*Euler's constant*)

$X_{vi} = 1, 0$ (*1 if correct answer; 0 if wrong*)

$-\infty < \beta_v < +\infty$ and $-\infty < D_i < +\infty$.

For example:

Suppose the ability of person v is $\beta=3$, while the difficulty level of item i is $D=1$. Then, the probability of getting a correct answer is .8808.

$$P_{vi} = \frac{e^{(3-1)}}{1 + e^{(3-1)}} = \frac{e^{(2)}}{1 + e^{(2)}} = .8808$$

Suppose the same person will attempt to answer an item with $D=2$. Then, the probability of getting a correct answer is .7311.

$$P_{vi} = \frac{e^{(3-2)}}{1 + e^{(3-2)}} = \frac{e^{(1)}}{1 + e^{(1)}} = .7311$$

Probabilities (P_{vi}) can be easily computed in Excel using the formula:

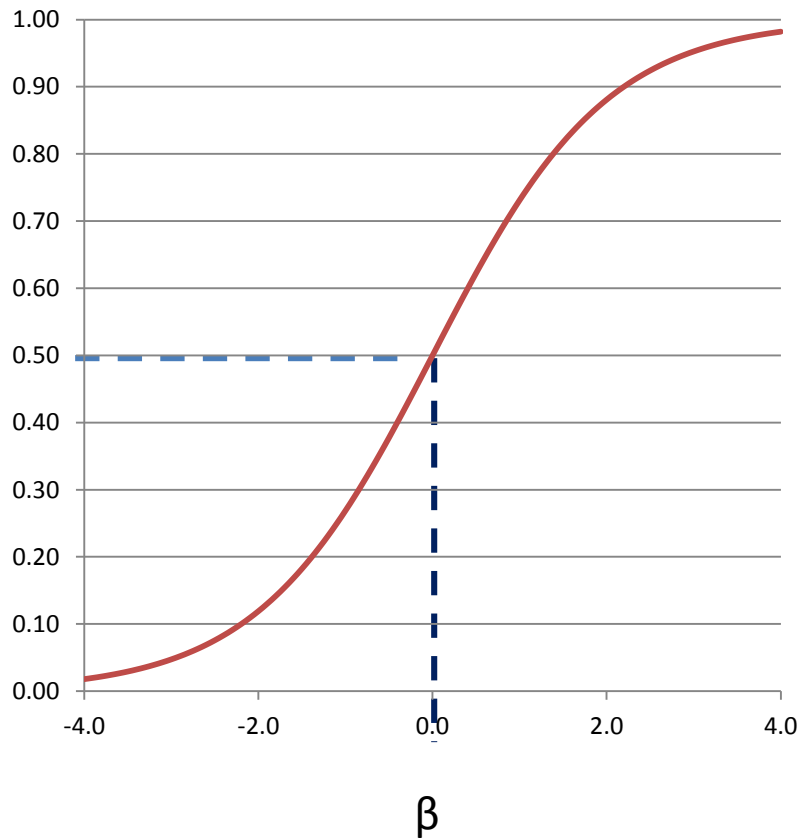
$$=Exp(B-D)/(1+EXP(B-D))$$

Excel outputs below:

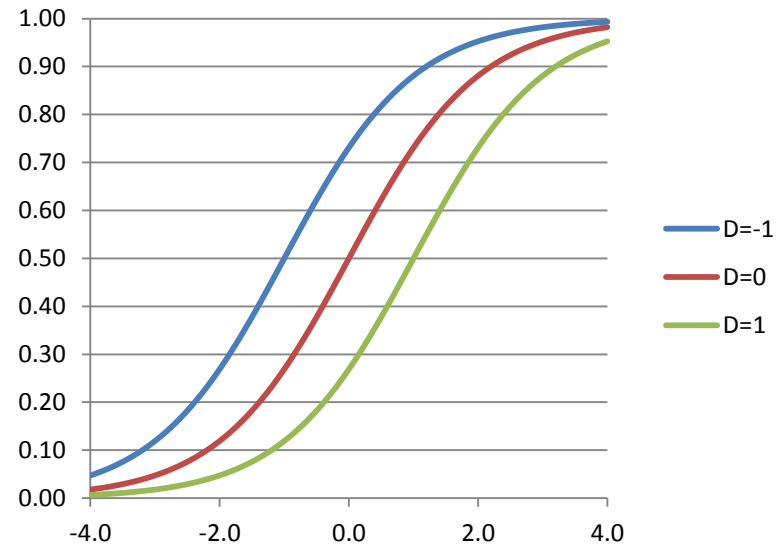
| | Item1 | Item2 | Item3 |
|----------|-------------|------------|------------|
| B | D=-1 | D=0 | D=1 |
| 4.0 | 0.9933 | 0.9820 | 0.9526 |
| 3.9 | 0.9926 | 0.9802 | 0.9478 |
| 3.8 | 0.9918 | 0.9781 | 0.9427 |
| 3.7 | 0.9910 | 0.9759 | 0.9370 |
| 3.6 | 0.9900 | 0.9734 | 0.9309 |
| 3.5 | 0.9890 | 0.9707 | 0.9241 |
| 3.4 | 0.9879 | 0.9677 | 0.9168 |
| 3.3 | 0.9866 | 0.9644 | 0.9089 |
| 3.2 | 0.9852 | 0.9608 | 0.9002 |
| 3.1 | 0.9837 | 0.9569 | 0.8909 |
| 3.0 | 0.9820 | 0.9526 | 0.8808 |
| 2.9 | 0.9802 | 0.9478 | 0.8699 |
| 2.8 | 0.9781 | 0.9427 | 0.8581 |

[Excelfile.xls](#)

Item Characteristic Curves (ICCs)



Items with different difficulty levels



Notes:

$$\text{For } \beta=D, \quad P_{vi} = \frac{e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}} = \frac{e^0}{1 + e^0} = .5$$

The probability of success is .5 if **person ability** matches the **item difficulty**.

$$\text{For } \beta > D, \quad P_{vi} > .5$$

The probability of success is greater than .5 if **person ability** is higher than the **item difficulty**.

$$\text{For } \beta < D, \quad P_{vi} < .5$$

The probability of success is lesser than .5 if the **person ability** is lesser than the **item difficulty**.

Odds for Success

Consider Person v:

Person v's odds of correctly getting a correct answer on item i, given his ability β_v and item difficulty D_i is defined as:

$$\frac{P_{vi}}{1 - P_{vi}} = \frac{\frac{e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}}}{1 - \frac{e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}}} = \frac{\frac{e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}}}{\frac{1 + e^{(\beta_v - D_i)} - e^{(\beta_v - D_i)}}{1 + e^{(\beta_v - D_i)}}} = e^{(\beta_v - D_i)}$$

Taking Logarithm:

$$\text{Log}_e \left(\frac{P_{vi}}{1 - P_{vi}} \right) = \text{Log}_e \left(e^{(\beta_v - D_i)} \right)$$

$$\text{Log}_e \left(\frac{P_{vi}}{1 - P_{vi}} \right) = \beta_v - D_i$$

Person v's log odds of correctly getting a correct answer on item i , given his ability β_v and item difficulty D_i is defined as:

$$\text{Log}_e \left(\frac{P_{vi}}{1 - P_{vi}} \right) = \beta_v - D_i$$

Consider another Person m:

Similarly, Person m 's log odds of getting a correct answer on the same item, given his ability β_m and item difficulty D_i :

$$\text{Log}_e \left(\frac{P_{mi}}{1 - P_{mi}} \right) = \beta_m - D_i$$

Comparing the abilities of Persons v and Person m:

To compare, we subtract the logarithm of odds:

$$(\beta_v - D_i) - (\beta_m - D_i) = \text{Log}_e \left(\frac{P_{vi}}{1 - P_{vi}} \right) - \text{Log}_e \left(\frac{P_{mi}}{1 - P_{mi}} \right)$$

$$\beta_v - \beta_m = \text{Log}_e \left(\frac{P_{vi}}{1 - P_{vi}} \right) - \text{Log}_e \left(\frac{P_{mi}}{1 - P_{mi}} \right)$$

Notice that the difference in the abilities of Person v and Person m does not involve D_i at all. This means that comparison on the person abilities does not depend on which particular item is used and so comparison is **“ITEM-FREE”**.

An analogous argument leads to **“PERSON-FREE”** comparisons of item difficulties.

Estimation of Rasch Parameters [1]

Winsteps implements the following methods of estimating Rasch parameters:

- JMLE (*Joint Maximum Likelihood Estimation by Wright and Panchapakesan*),
- PROX (*Normal Approximation Algorithm devised by Cohen (1979)*).

Estimation of Rasch Parameters [2]

Rasch measures are obtained by iterating through the data.

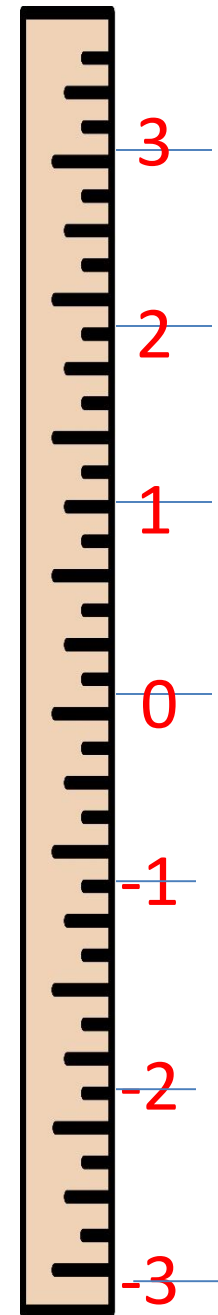
- **STEP 1:** Initially all unanchored parameter estimates (measures) are set to zero.
- **STEP 2:** Then the PROX method is employed to obtain rough estimates. Each iteration through the data improves the PROX estimates until they are usefully good.
- **STEP 3:** Then those PROX estimates are the initial estimates for JMLE which fine-tunes them, again by iterating through the data, in order to obtain the final JMLE estimates. The iterative process ceases when the convergence criteria are met.

Convergence Criteria

- Largest Logit change
(default of Winsteps: LCONV=.0001 logits)
- Largest Score Residual
(default of Winsteps: RCONV=.01 score points)
- MJMLE= 0 ; unlimited JMLE iterations

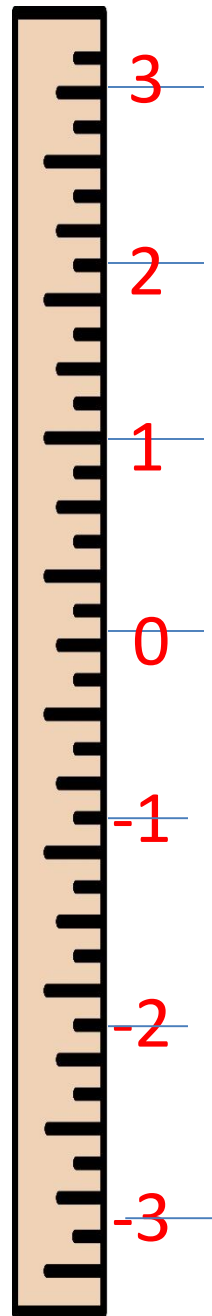
Rasch “Yardstick” [1]

- Rasch model creates a “yardstick” that can be used to measure both **Person Ability** and **Item Difficulty**.
- **The values in the yardstick are logits, which range between $-\infty$ and $+\infty$. But for application purposes, they range between -4 and +4 or between -3 and +3.**

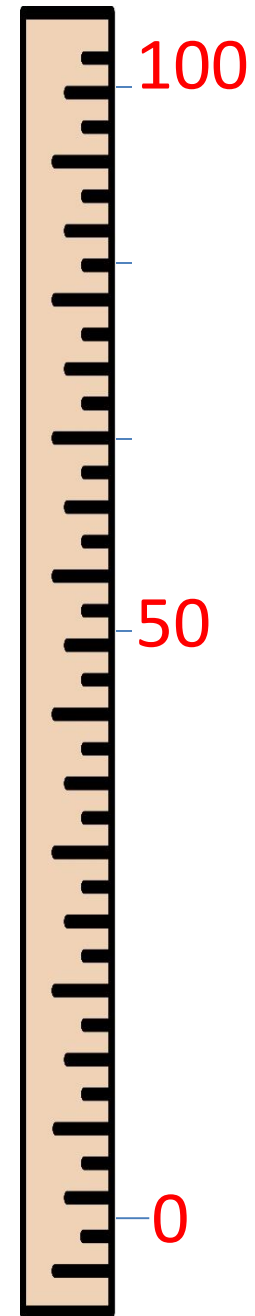


Rasch “Yardstick” [2]

- The logits can be transformed so that the range would be understandable by non-technical users.
- For example, the logits can be transformed so that the values would fall between 0 and 100.

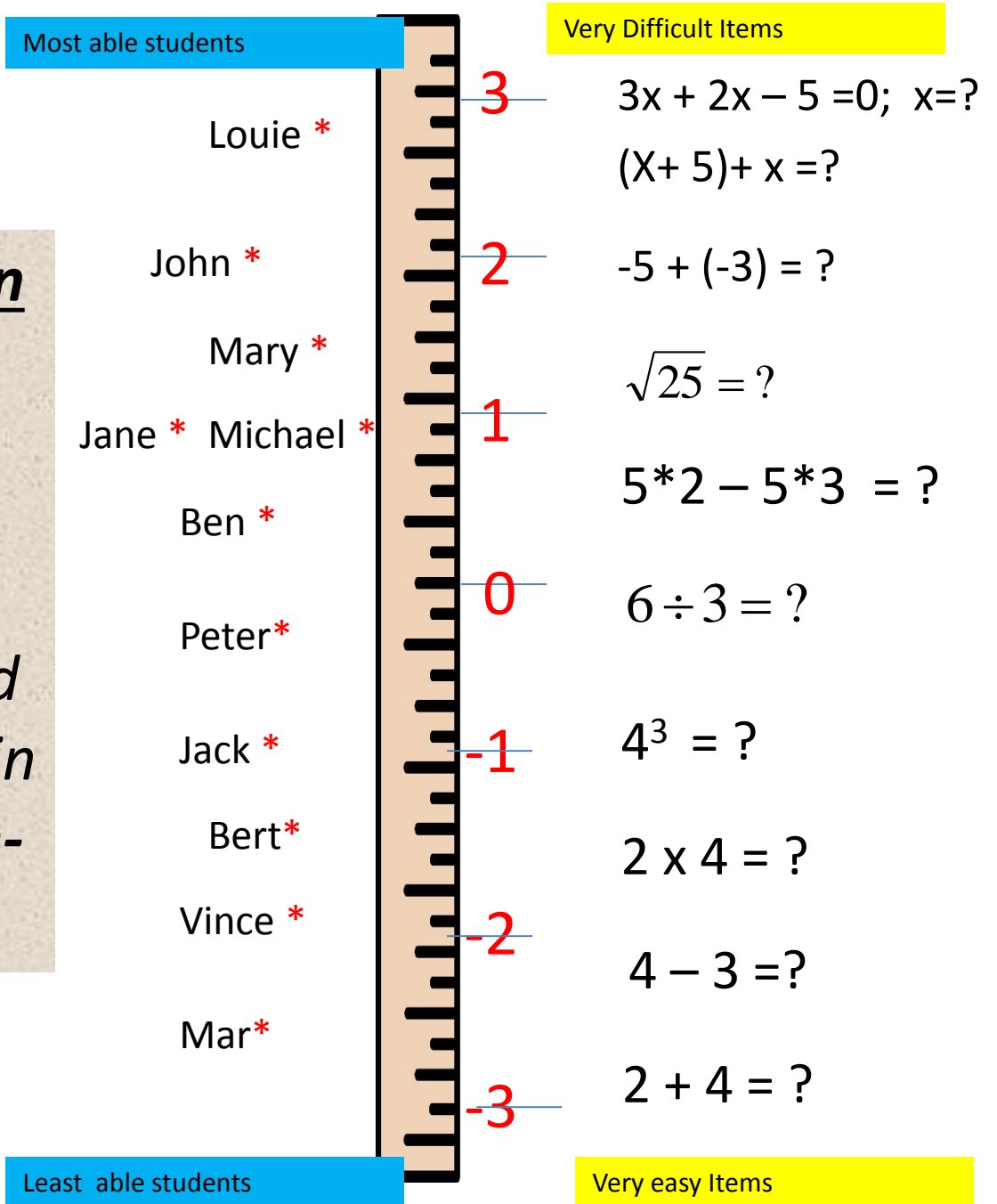


Transformed Version



Person-Item Map [1]

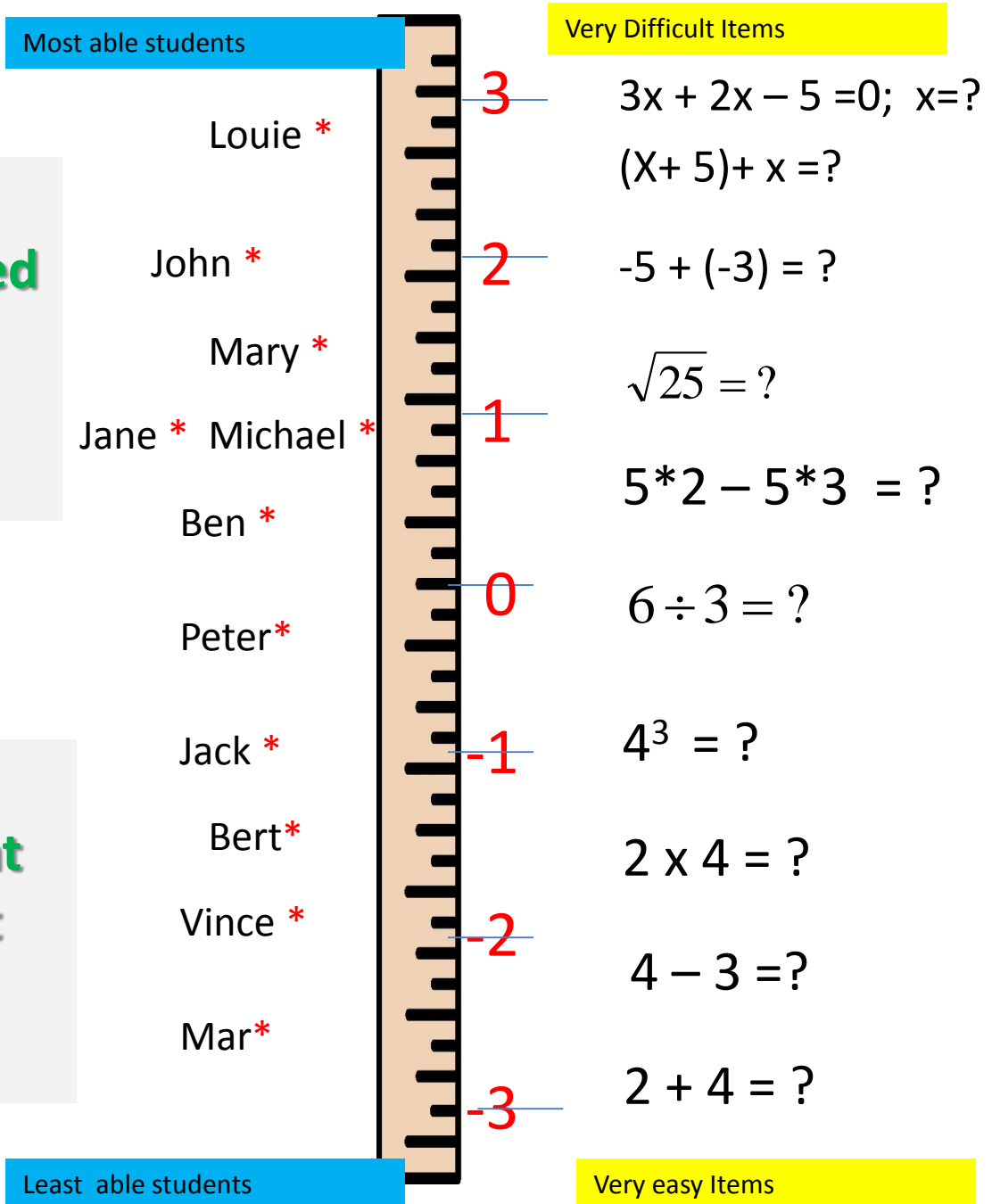
Because Person ability and item difficulty are measured using a common “yardstick, then both person and items can be placed in a map, called **Person-Item map**.



Person-Item Map [2]

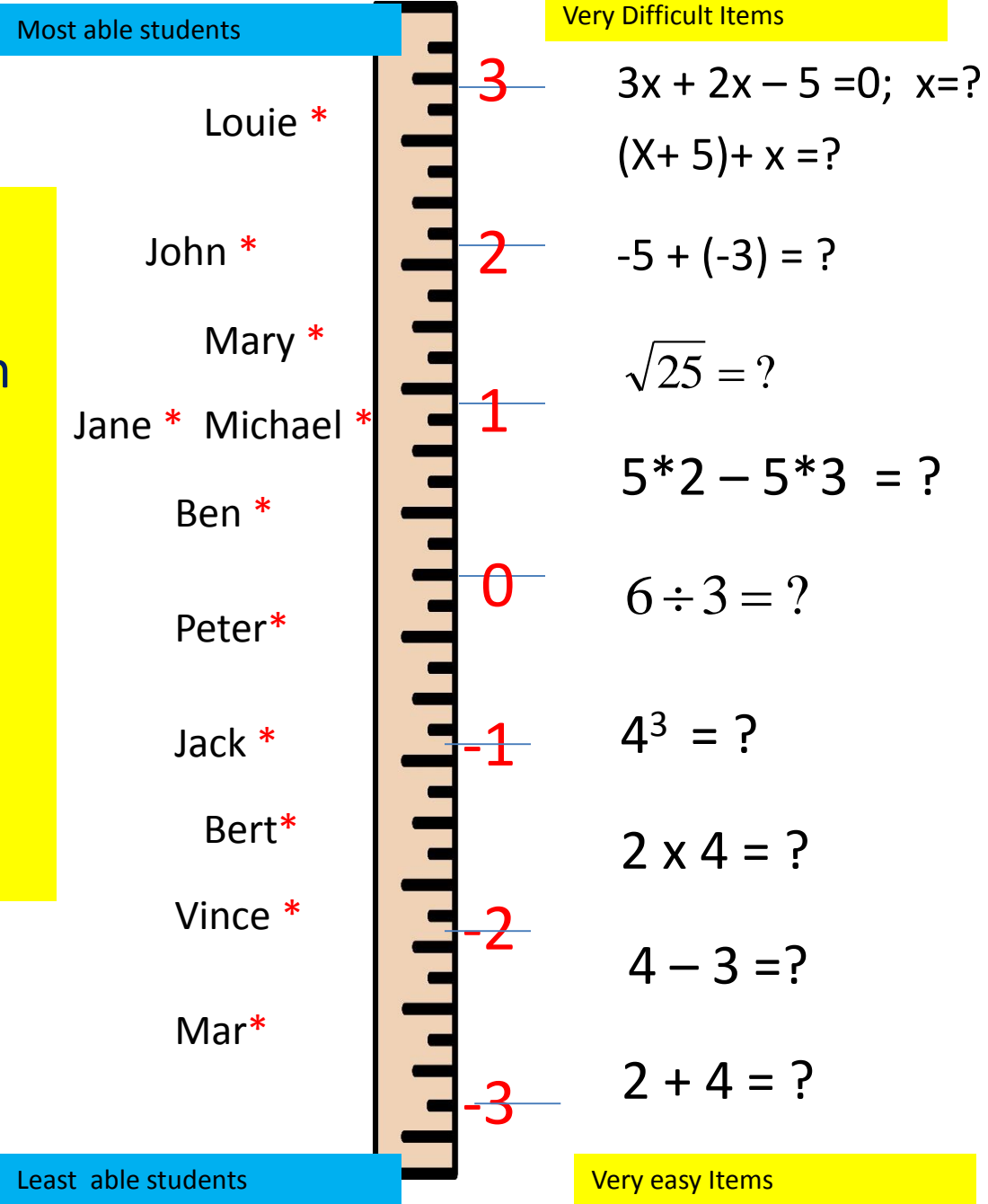
- Items are hierarchically arranged from very easy (bottom) up to very difficult (top).

- Also, least able students are placed at the bottom and most able students at the top.



Person-Item Map [3]

The nice with Rasch modeling is that we can determine which students are able to answer correctly which items. Or, we can determine which items can be answered correctly by which students.

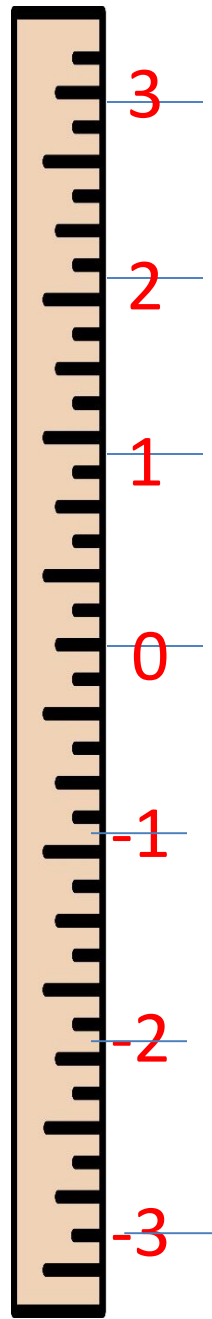


Things to consider in constructing a test using Rasch Model

- Validity: Construct Validity, Fit validity
- Reliability

Construct Validity:

Does the item difficulty hierarchy make sense?



3x + 2x - 5 = 0; x=?
(X+ 5)+ x =?

-5 + (-3) = ?

$\sqrt{25} = ?$

5*2 - 5*3 = ?

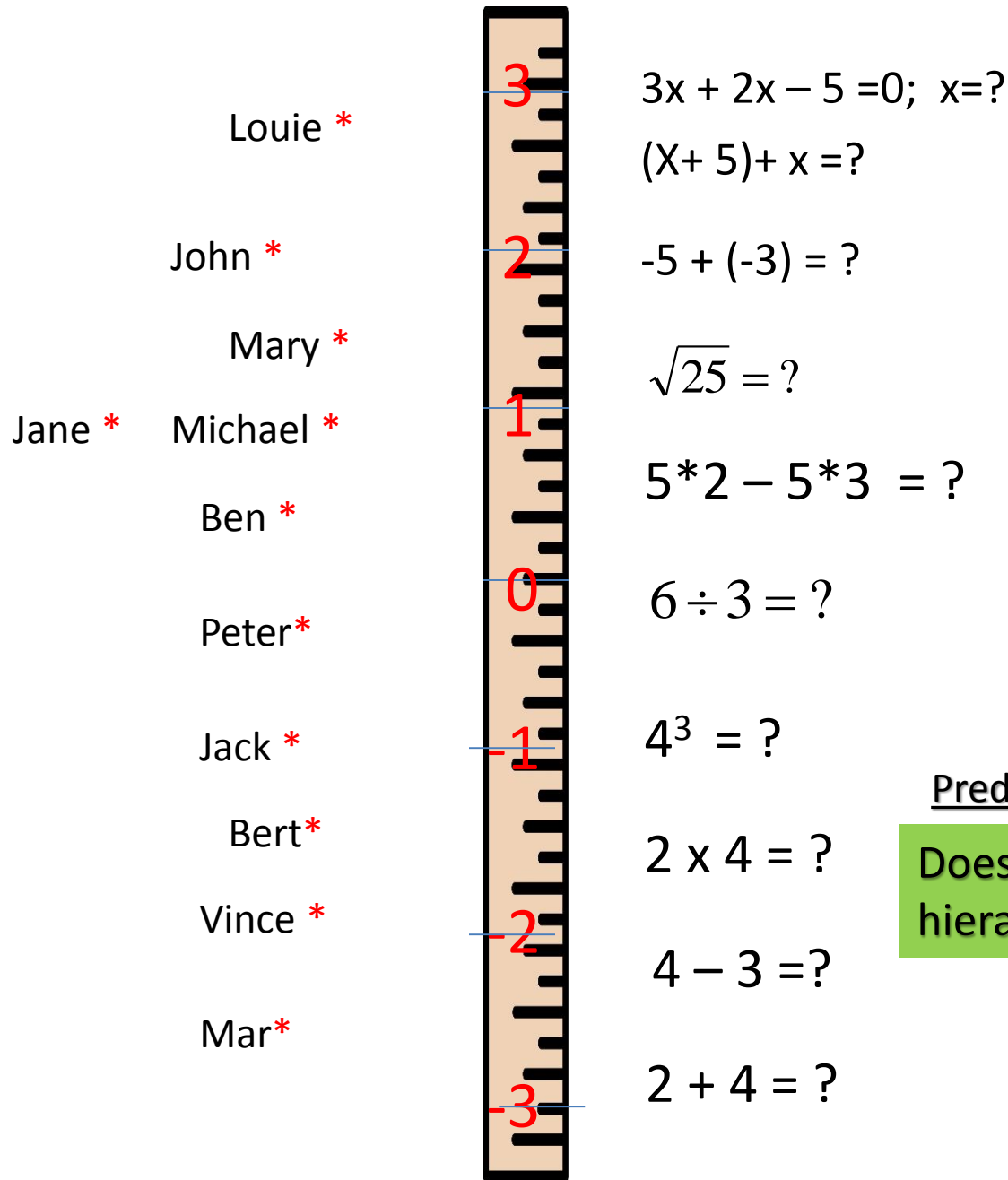
6 ÷ 3 = ?

4³ = ?

2 x 4 = ?

4 - 3 = ?

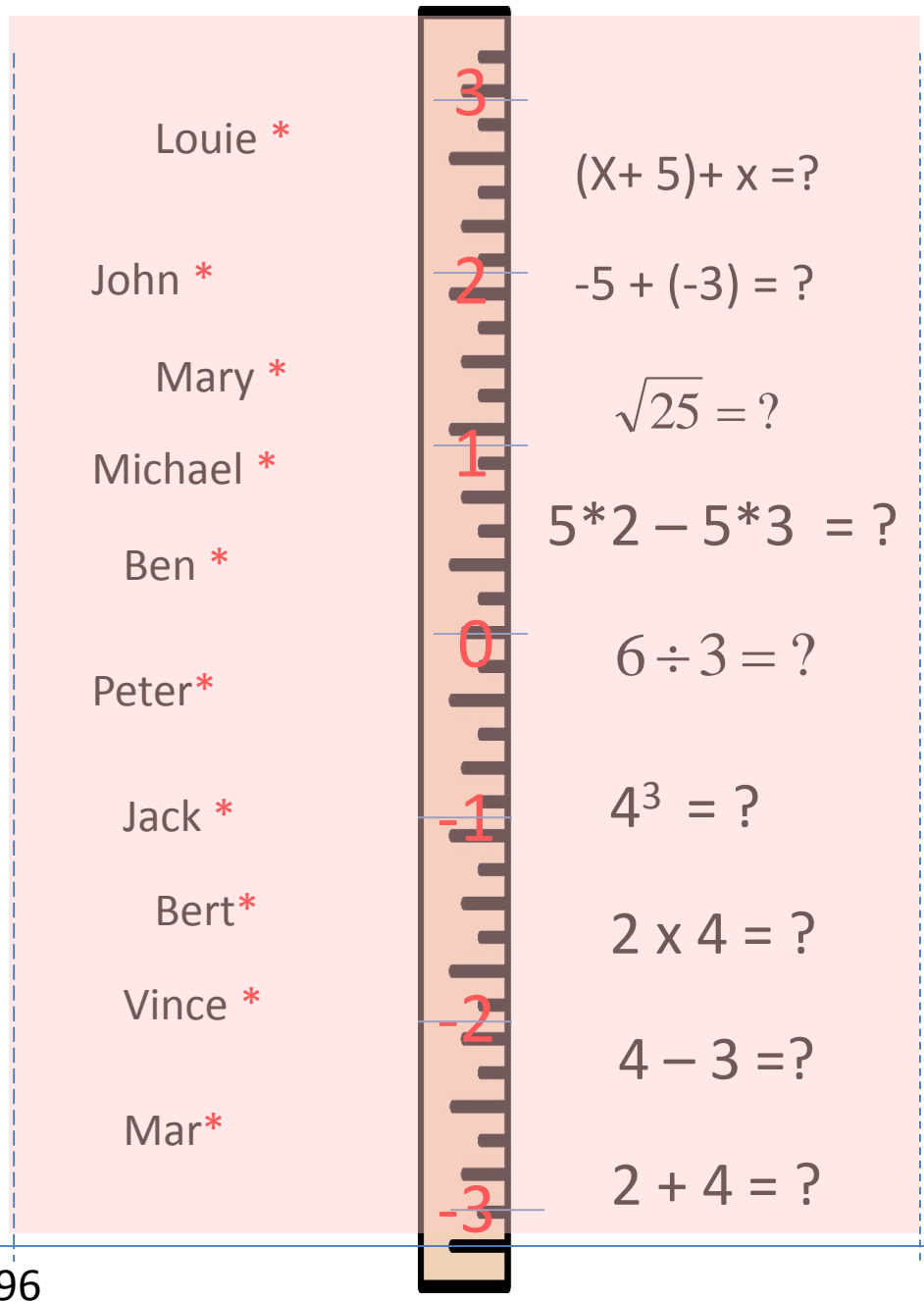
2 + 4 = ?



Predictive Validity:

Does the person ability hierarchy make sense?

James*



$S = v \times t$

Fit Validity:
Do data fit the rasch model usefully well for the purpose of measurement?

Z = -1.96

Z = 1.96

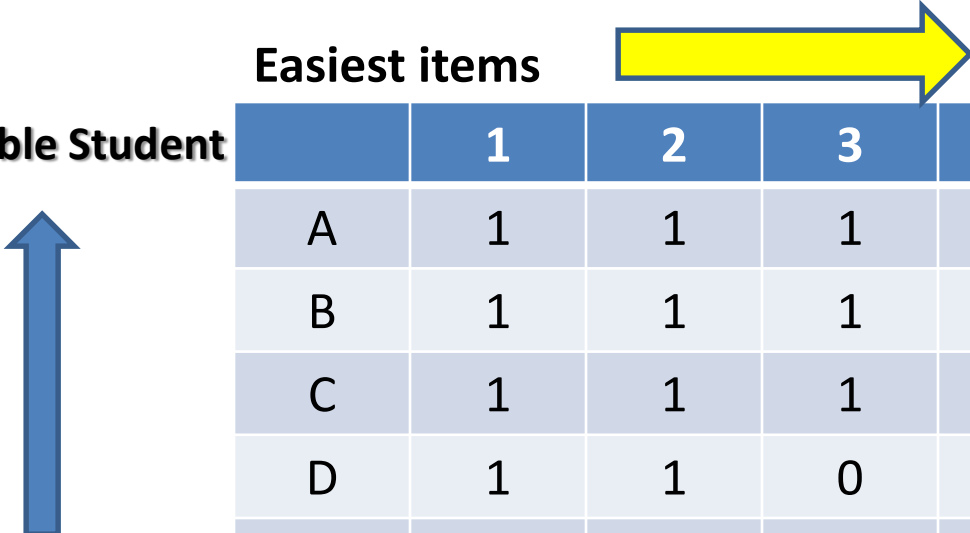
Concept of Model and Data Fit

- Major concern of Rasch Modeling is its need for unidimensionality.
- Investigation of fit statistics determines whether the data are unidimensional in nature.
- Both **infit and outfit statistics** are evaluated to determine how **data-to-model** fit occurs for each item and person fit.

Infit and Outfit Statistics

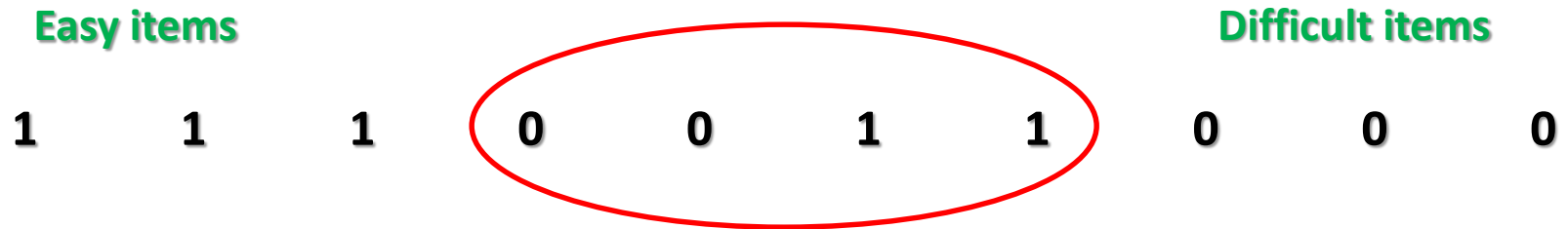
- Infit statistics are sensitive to the inlier pattern of observations.
- Outfit statistics are sensitive to outlier observations.

Idealized Guttman Scale (Gutman, 1944)



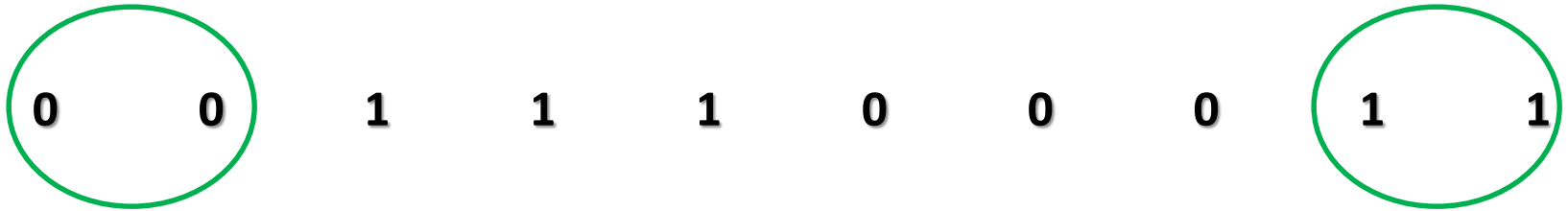
| | Easiest items | | | Hardest Items | | |
|--------------------|---------------|---|---|---------------|---|---|
| Most Able Student | 1 | 2 | 3 | 4 | 5 | 6 |
| A | 1 | 1 | 1 | 1 | 1 | 0 |
| B | 1 | 1 | 1 | 1 | 0 | 0 |
| C | 1 | 1 | 1 | 0 | 0 | 0 |
| D | 1 | 1 | 0 | 0 | 0 | 0 |
| Least Able Student | E | 1 | 0 | 0 | 0 | 0 |

Data with large infit statistics



Larger infit statistics because the 1's occurring in the middle-right section of the continuum and the 0's appearing in the middle-left section of the continuum are unexpected.

Data with large outfit statistics



Larger outfit statistics because observations at the extreme ends of the continuum are unexpected.

Fit Statistics as Indicator of Validity

Data adequately fitting the model is a key indicator of validity.

- Removal of misfitting persons and items that grossly misfit the model's expectation is acceptable.

[Removal of misfitting persons and items improves the precision of the measures produced.]

Formulas: *Fit Statistics*

- Outfit Mean Square: outlier-sensitive fit statistic. This is based on the conventional chi-square statistic.

$$\text{Outfit Mean Square} = \text{average} [(\text{standardized residuals}^2)] = \text{chi-square/d.f.}$$

- Infit Mean Square: inlier-pattern-sensitive fit statistic. This is based on the chi-square statistic with each observation weighted by its statistical information (model variance).

$$\text{Infit Mean Square} = \text{average} [(\text{standardized residuals}^2) * \text{information}]$$

- Z-Standardized: statistical significance (probability) of the chi-square (mean-square) statistics occurring by chance when the data fit the Rasch model. The values reported are unit-normal deviates.

| ZSTD probabilities: two-sided unit-normal deviates | |
|---|----------|
| 1.00 | p= .317 |
| 1.96 | .050 |
| 2.00 | .045 |
| 3.00 | .0027 |
| 4.00 | .00006 |
| 5.00 | .0000006 |

Infit Mean Square and Outfit Mean Square: Rule of Thumbs

| | |
|------------------|---|
| | |
| > 2.0 | Distorts or degrades the measurement system. |
| 1.5 - 2.0 | Unproductive for construction of measurement, but not degrading. |
| 0.5 - 1.5 | Productive for measurement. |
| <0.5 | Less productive for measurement, but not degrading. May produce misleadingly good reliabilities and separations. |

Note: **1** = Expected Value (perfect fit)

Reliability

- Reliability means reproducible of relative measure location.
- “High item reliability” means that there is a high probability that items estimated with high measures actually do have higher measures than items estimated with low measures.
- “High person reliability” means that there is a high probability that persons estimated with high measures actually do have higher measures than persons estimated with low measures.

Person Reliability

Person reliability depends chiefly on:

- Sample ability variance. **Wider ability range = higher person reliability.**
- Length of test. **Longer test = higher person reliability**
- Sample-item targeting. **Better targeting = higher person reliability**

How to increase person reliability?

- test persons with more extreme abilities (high and low)
- lengthen the test.

Person Reliability is independent of sample size. It is largely uninfluenced by model fit.

Item Reliability

Item reliability depends chiefly on

- **Item difficulty variance. Wide difficulty range = high item reliability**
- **Person sample size. Large sample = high item reliability**

How to increase item reliability?

- test more people.

Item Reliability is independent of test length. It is largely uninfluenced by model fit.

Sample Size Requirements

- Rasch is the same as any other statistical analysis with a small sample:
 - Less precise estimates (*bigger standard errors*)
 - Less powerful fit analysis
 - Less robust estimates
- Very small sample (*say, $n=2$ or 3 examinees*) provides a very unstable results, while very large sample (*say, $n=2000$ or 3000*) provides a very precise results. However, large sample is too expensive and time-consuming. So, how big a sample is necessary?

- Linacre (1994) provides the following sample size guidelines:

| Item Calibrations stable within | Confidence | Minimum sample size range (best to poor targeting) | Size for most purposes |
|---------------------------------|--------------|--|---------------------------------|
| ± 1 logit | 95% | 16 † -- 36 | 30 (minimum for dichotomies) |
| ± 1 logit | 99% | 27 † -- 61 | 50 (minimum for polytomies) |
| $\pm \frac{1}{2}$ logit | 95% | 64 -- 144 | 100 |
| $\pm \frac{1}{2}$ logit | 99% | 108 -- 243 | 150 |
| Definitive or High Stakes | 99%+ (Items) | 250 -- 20*test length | 250 |
| Adverse Circumstances | Robust | 450 upwards | 500 |

Reference:

Linacre JM. (1994). Sample Size and Item Calibration Stability. Rasch Measurement Transactions, 7:4 p.328

Some Features of Rasch Model

1. Examinee performance on an unadministered item can be predicted.
2. Item and ability parameters can be estimated
3. Item parameter estimates are independent of the group of examinees who took the test
4. Examinee ability estimates are independent of the group of test items administered
5. Precision of ability estimates is known

Rasch versus 1PL IRT

| Aspect | Rasch | IRT: One-parameter Logistic Model |
|-----------------------------------|---|--|
| Symbol | Rasch | 1PL IRT, also 1PL |
| For Practical purposes | When each individual in the person sample is parameterized for item estimation, it is Rasch. | When the person sample is parameterized by a mean and standard deviation for item estimation, it is 1PL IRT. |
| <i>Motivation</i> | <p>Prescriptive: Distribution-free person ability estimates and distribution-free item difficulty estimates on an additive latent variable</p> | <p>Descriptive: Computationally simpler approximation to the Normal Ogive Model of L.L. Thurstone, D.N. Lawley, F.M. Lord</p> |
| Formulation: Exponential Form | $P_n = \frac{e^{a_n - D_i}}{1 + e^{a_n - D_i}}$ | $P_i(\theta) = \frac{e^{1.7(\theta - b_i)}}{1 + e^{1.7(\theta - b_i)}}$ |
| Formulation: Logit-linear form | $\log_e \left(\frac{P_n}{1 - P_n} \right) = B_n - D_i$ | $\log_e \left(\frac{P_i(\theta)}{1 - P_i(\theta)} \right) = 1.7(\theta - b_i)$ |

| Aspect | Rasch | IRT: One-parameter Logistic Model |
|--|--|--|
| <i>Students/persons</i> | Person n of ability B_n in logits | Normally-distributed person sample of ability distribution θ , conceptualized as $N(0,1)$, in probits; persons are incidental parameters |
| <i>Items, multiple-choice questions, etc.; items are structural parameters</i> | Item i of difficulty D_i in logits | Item i of difficulty b_i (the "one parameter") in probits |
| <i>Nature of binary data</i> | 1 = "correct" 0 = "wrong" | 1 = "correct" 0 = "wrong" |
| <i>Probability of binary data</i> | P_{ni} = probability that person n correctly answered item i | $P_i(\theta)$ = overall probability of "correct" by person distribution θ on item i |

| Aspect | Rasch | IRT: One-parameter Logistic Model |
|---|---|--|
| <i>Local origin of scale: zero of parameter estimates</i> | Average item difficulty, or difficulty of specified item. (Criterion-referenced) | Average person ability. (Norm-referenced) |
| <i>Item discrimination</i> | Item characteristic curves (ICCs) modeled to be parallel with a slope of 1 (the natural logistic ogive) | ICCs modeled to be parallel with a slope of 1.7 (approximating the slope of the cumulative normal ogive) |
| <i>Fit evaluation</i> | <i>Fit of the data to the model</i> Local, one parameter at a time | <i>Fit of the model to the data</i> Global, accept or reject the model |
| <i>Data-model mismatch</i> | Defective data do not support parameter separability in an additive framework. Consider editing the data. | Defective model does not adequately describe the data. Consider adding discrimination (2-PL), lower asymptote (guessability, 3-PL) parameters. |

| Aspect | Rasch | IRT: One-parameter Logistic Model |
|---|--|---|
| <i>Minimum reasonable sample size</i> | 30 Linacre (1994) | 200 (Downing 2003) |
| First conspicuous appearance | Rasch, Georg. (1960) Probabilistic models for some intelligence and attainment tests. Copenhagen: Danish Institute for Educational Research. | Birnbaum, Allan. (1968). Some latent trait models. In F.M. Lord & M.R. Novick, (Eds.), Statistical theories of mental test scores. Reading, MA: Addison-Wesley. |
| First conspicuous advocate | Benjamin D. Wright, University of Chicago | Frederic M. Lord, Educational Testing Service |
| Widely-authoritative currently-active proponent | David Andrich, Univ. of Western Australia, Perth, Australia | Ronald Hambleton, University of Massachusetts |

Source (of the comparison):

Linacre J.M. (2005). Rasch dichotomous model vs. One-parameter Logistic Model. Rasch Measurement Transactions, 19:3, 1032

Demonstration & Computer Hands-on

Rasch Analysis using Winsteps/Ministep

Download data and installers from this link:

https://www.dropbox.com/sh/vkxfosv6c630ftm/AAB-A2NqO2QG44J_Wa8CuvAsa?dl=0

Download Free Ministep

www.winsteps.com/ministep.htm

Free - Freeware - Rasch measurement software ...



MINISTEP

Evaluation, Student and Demonstration (Demo) Version of WINSTEPS

Free by Download

MINISTEP is a reduced version of [WINSTEPS](#). It has complete WINSTEPS functions but is limited to 25 items and 75 persons. You can evaluate and use it without charge or time-limit. Of course, we hope you will want to purchase the full version.

Get started with [Winsteps Tutorial PDFs](#)

- | | |
|--|---|
| Ministep: Standard Download  free download NOW! | Ministep: Secure Download from https://mmm1425.sanjose14-verio.com  free download NOW! |
|--|---|
- "Save as" "c:\windows\desktop\MinistepInstall.exe"
- Click on "**MinistepInstall.exe**" on your desktop to install MINISTEP, the free version. MINISTEP will start automatically.
- Delete "MinistepInstall.exe"
- Click on "Ministep" icon on desktop to run MINISTEP
- [Installation problems?](#)

- Ministep in a zip file: [MinistepInstall.zip](#)

MINISTEP:

- 25 items
- 75 persons
- WINSTEPS
 - 30,000 items
 - 10,000 persons

Example 1- Math Curriculum Test

| Persons | Items | | | | | | | | | | | | Raw Score |
|---------|----------|----------|----------|----------|----------|----------|----------|----------|----------|----------|----------|----------|-----------|
| | <i>a</i> | <i>b</i> | <i>c</i> | <i>d</i> | <i>e</i> | <i>f</i> | <i>g</i> | <i>h</i> | <i>i</i> | <i>j</i> | <i>k</i> | <i>l</i> | |
| A | ✓ | ✓ | ✓ | X | X | X | X | ✓ | ✓ | X | X | ✓ | 6 |
| B | ✓ | X | ✓ | ✓ | X | X | X | X | ✓ | X | X | X | 4 |
| C | ✓ | ✓ | ✓ | X | ✓ | X | X | ✓ | ✓ | ✓ | ✓ | ✓ | 9 |
| D | ✓ | X | ✓ | ✓ | X | X | X | X | ✓ | X | X | ✓ | 5 |
| E | X | ✓ | ✓ | X | X | ✓ | X | ✓ | ✓ | ✓ | ✓ | ✓ | 8 |
| F | ✓ | ✓ | ✓ | ✓ | ✓ | X | X | ✓ | ✓ | X | X | ✓ | 8 |
| G | ✓ | X | ✓ | X | X | ✓ | X | X | ✓ | X | ✓ | ✓ | 6 |
| H | ✓ | X | ✓ | X | X | X | X | X | X | X | X | ✓ | 3 |
| I | ✓ | ✓ | ✓ | ✓ | X | X | X | ✓ | ✓ | X | X | ✓ | 7 |
| J | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | X | ✓ | X | 10 |
| K | ✓ | X | ✓ | X | X | ✓ | X | X | ✓ | X | ✓ | ✓ | 6 |
| L | X | ✓ | ✓ | X | X | ✓ | X | ✓ | ✓ | ✓ | ✓ | ✓ | 8 |
| M | X | X | X | X | X | X | X | X | X | X | X | X | 0 |
| N | ✓ | ✓ | ✓ | ✓ | X | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | 11 |

12 items
14 persons

Q & A

References

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Thank you!

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